

# Threshold Policies for Single-Resource Reservation Systems

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## 1 Introduction

Requests for a resource arrive at rate  $\lambda$ , each request specifying a future time interval, called a *reservation interval*, to be booked for its use of the resource. The *advance notices* (delays before reservation intervals are to begin) are independent and drawn from a distribution  $A(z)$ . The durations of reservation intervals are sampled from the distribution  $B(z)$  and are independent of each other and the advance notices. We let  $A$  and  $B$  denote random variables with the distributions  $A(z)$  and  $B(z)$  (the functional notation will always allow one to distinguish between our two uses of the symbols  $A$  and  $B$ ).

The following greedy reservation policy was analyzed in [3]: A request is immediately accepted (booked) if and only if the resource will be available throughout its reservation interval, i.e., the resource has not already been reserved for a time period overlapping the requested reservation interval. In [3], the authors compute an efficiency measure, called the *reservation probability*, which is the fraction of time the resource is in use.

This paper studies the reservation probability for a more general greedy policy of threshold type that is defined by two parameters  $s$  and  $\tau$ . If a request has an advance notice less than  $s$  or a duration exceeding  $\tau$ , then the threshold policy makes an attempt to book it under the greedy rule; otherwise, it is rejected even if it could have been accommodated. Our main result is an expression for the asymptotic reservation probability as  $s \rightarrow \infty$  and the advance-notice distribution becomes progressively more spread out.

The above result relates asymptotics of reservation policies to asymptotics of interval packing policies, a connection first studied in [3]. In the interval packing problem [1], intervals arrive randomly in  $\mathbf{R}_+^2$  according to a Poisson process in the two dimensions representing arrival times  $t$  and the left endpoints of the arriving intervals. Interval lengths are i.i.d., and since we will map them to reservation intervals, we let their distribution also be denoted by  $B(z)$ . The intensity is 1, i.e., an average of one interval arrives per unit time per unit distance. For a given  $x > 0$ , an arriving interval is packed (or accepted) in the ‘containing’ interval  $[0, x]$  under the greedy algorithm if and only if it is a subinterval of  $[0, x]$  and it does not overlap an interval already accepted. The problem is to find, or at least estimate, the function  $K(t, x)$ , which is the expected total length of the intervals accepted by the greedy policy during  $[0, t]$ , assuming that none has yet been accepted by time 0 ( $[0, x]$  is initially empty).

Estimates of  $K(t, x)$  were obtained in [3] from its Laplace transform  $\mathcal{K}(t, u)$ ; these results are special cases of the corresponding results for the *threshold* packing policy with parameters  $s, \tau$ . The threshold packing policy extends greedy interval packing much as we extended the greedy reservation policy: An interval is processed by the greedy packing algorithm if its length is at least  $\tau$  or if it arrives no sooner than  $s$ ; otherwise, it is rejected. The next section exhibits the Laplace transform of  $H_\tau(s, t, x)$ , the expected total length of the intervals accepted during  $[0, t]$ ,  $t \geq s$ , by the threshold packing policy with parameters  $s, \tau$ . Note that threshold packing reduces to simple greedy packing if  $\tau = 0$  or if  $s = 0$ . The formulas in the next section will verify that  $K(t, x) = H_0(t, t, x)$ .

As noted in [3], there are many potential applications covered by models like ours. However, relatively new applications in existing and proposed communication systems, e.g., teleconferencing and video-on-demand systems, have given a fresh impetus to research on reservation systems. Previous work in the communications field is quite recent and focuses more on engineering problems than mathematical foundations; past research has dealt with the implementation issues of incorporating distributed advance-notice reservation protocols in current networks, and with the algorithmic issues concerned with well utilized resources in reservation systems (see [3, 4, 5] for many references). For the analysis of mathematical models different from our own, see the work of Virtamo [5] and Greenberg, Srikant, and Whitt [4].

## 2 Threshold interval packing

Let  $b_\tau := \mathbf{E}B \cdot \mathbf{1}(B > \tau)$ ,  $b_\tau^{(2)} := \mathbf{E}B^2 \cdot \mathbf{1}(B > \tau)$ , and  $p = p_\tau := \mathbf{P}(B > \tau)$ . Denote by  $L_\tau(s, t, x)$  the total length of the intervals packed at time  $t \geq s$  in  $[0, x]$ , and let  $H_\tau(s, t, x) := \mathbf{E}L_\tau(s, t, x)$ .

We compute the rate of change of  $H_\tau(s, t, x)$  with respect to  $s$  for  $x \geq d$  by expressing  $H_\tau(s + \Delta s, t, x)$  in terms of  $H_\tau(s, t, x)$  and the events occurring in the time interval  $[0, \Delta s]$ . On doing so, rearranging the recurrence and taking the limit  $\Delta \rightarrow 0$ , we obtain the integro-differential equation

$$\begin{aligned} \frac{\partial H_\tau(s, t, x)}{\partial s} &= -(px - b_\tau)H_\tau(s, t, x) + b_\tau x - b_\tau^{(2)} \\ &\quad + 2 \int_\tau^d dB(z) \int_0^{x-z} H_\tau(s, t, y) dy. \quad (1) \end{aligned}$$

Note the boundary condition  $K(t, x) = H_0(t, t, x)$ , the known result for the simple greedy rule. Note also that the effect on (1) of putting  $\tau = 0$  is confined to constants depending only on  $\tau$ . Thus, exactly the same analysis for  $K$  in [3] can be applied to  $H_\tau$  here. We introduce the transforms  $\mathcal{H}_\tau(s, t, u) := \int_d^\infty e^{-ux} H_\tau(s, t, x) dx$ ,  $\mathcal{B}_\tau(u) := \int_\tau^d e^{-uz} dB(z)$ , and then transform (1) to obtain a pde whose solution is readily found to be

$$\mathcal{H}_\tau(s, t, u) = \frac{1}{u^2} \int_0^s \mathcal{C}_\tau(s-z, t, u+pz) \mathcal{G}_\tau(u, pz) dz \quad (2)$$

$$+ \left( \frac{u+ps}{u} \right)^2 \mathcal{K}(t-s, u+ps) \mathcal{G}_\tau(u, ps), \quad x \geq d,$$

where  $\mathcal{G}_\tau(u, v) := \exp\left(b_\tau v - 2 \int_u^{u+v} \frac{1-B_\tau(y)}{y} dy\right)$ , and where  $\mathcal{C}_\tau(s, t, u)$  is the transform of a well-behaved function easily computed by a recursive procedure. (See the expanded version of the paper [2] for details.) As illustrated below, explicit expressions are available for simple distributions  $B(\cdot)$ .

**Estimates of  $H_\tau$ .** Leading-term asymptotics in  $x$  are given by the following result, in which

$$\alpha_\tau(s, t) := \int_0^s \mathcal{C}_\tau(s-z, t, pz) \mathcal{G}_\tau(pz) dz$$

$$+ \mathcal{K}(t, ps) (ps)^2 \mathcal{G}_\tau(ps). \quad (3)$$

**Theorem 1** For any  $s > 0$ ,  $t \geq 0$ ,  $H(s, t, x) \sim \alpha_\tau(s, t)x$  as  $x \rightarrow \infty$ .

**Proof sketch:** One first verifies that the functions  $\mathcal{C}_\tau$ ,  $\mathcal{G}_\tau$ , and  $\mathcal{K}$  are such that  $\mathcal{H}_\tau$  in (2) satisfies  $\mathcal{H}_\tau(s, t, u) \sim \alpha_\tau(s, t)/u^2$  as  $u \downarrow 0$ . An application of Karamata's Tauberian theorem and routine manipulations then proves the theorem. ■

**Example 1:** Consider the case where interval lengths have only the values 1 or 2, with probabilities  $\mathbf{P}(B=2) = 1 - \mathbf{P}(B=1) = q$ . Computations give

$$\mathcal{G}(u, v) = \exp\left(2q - 2 \int_u^{u+v} \frac{1-qe^{-2y}}{y} dy\right),$$

and

$$\mathcal{C}_\tau = 2qe^{-2u} \left( 1 + u \left[ \frac{e^{-u} - e^{-2u}}{u} - \frac{e^{-u} - e^{-2u-(1-p)t}}{u + (1-p)t} \right] \right),$$

from which we obtain  $\mathcal{K}(t, u)$  and hence an explicit integral formula for  $\alpha(s, t)$  in (3) that we can evaluate numerically.

For  $t = 7$ ,  $p = 0.4$ , we used Mathematica to compute the curve for  $\alpha_\tau(s, 10-s)$ . The threshold algorithm is strictly better than the greedy policy ( $\alpha_\tau(0, 10)$ ) for a large range of values of the threshold parameter  $s$ . ■

In order to obtain higher-order error terms, we use complex analysis and the Cauchy residue theorem to evaluate directly

the Laplace-transform inversion formula for  $\mathcal{H}_\tau$ . This gives the main result needed for the limit law of the next section,

**Theorem 2** For any fixed  $\xi, T > 0$  with  $\xi > 2T > 0$ , there exists a constant  $\gamma_\tau(t)$ , such that

$$\sup_{0 \leq t_1, t_2 \leq T} |H_\tau(t_1, t_2, x) - (\alpha_\tau(t)x + \gamma_\tau(t))| = O(e^{-\xi x}).$$

### 3 Advanced-notice limit law

Let the advance notice distribution be uniform on  $[0, a]$  and for a given  $\delta$ , ( $0 < \delta < 1$ ) consider the threshold reservation policy: If the advance notice of a request is in  $[\delta a, a]$  and the duration requested is less than a given  $\tau$ , then the request is rejected. Otherwise, it is processed according to the greedy reservation policy (i.e., it is accepted if does not overlap a time interval already reserved). Let  $P_a(\lambda, \tau, \delta)$  denote the reservation probability, i.e., the stationary probability that the resource is in use.

**Theorem 3** As the support of the advance notice distribution tends to infinity ( $a \rightarrow \infty$ ), we have the asymptotic reservation probability  $P_a(\lambda, \tau, \delta) \sim \alpha_\tau(\delta\lambda, \lambda)$ .

The proof follows the approach used in [3] to establish the analogous result for the simple greedy policy.

### References

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