Maximums on Trees

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Abstract

We study the minimal/endogenous solution R to the maximum recursion on weighted branching trees given by

$$R \stackrel{\mathcal{D}}{=} \left(\bigvee_{i=1}^{N} C_i R_i\right) \lor Q,$$

where $(Q, N, C_1, C_2, ...)$ is a random vector with $N \in \mathbb{N} \cup \{\infty\}$, P(|Q| > 0) > 0 and nonnegative weights $\{C_i\}$, and $\{R_i\}_{i\in\mathbb{N}}$ is a sequence of i.i.d. copies of R independent of $(Q, N, C_1, C_2, ...)$; $\stackrel{\mathcal{D}}{=}$ denotes equality in distribution. Furthermore, when Q > 0 this recursion can be transformed into its additive equivalent, which corresponds to the maximum of a branching random walk and is also known as a high-order Lindley equation. We show that, under natural conditions, the asymptotic behavior of R is power-law, i.e., $P(|R| > x) \sim Hx^{-\alpha}$, for some $\alpha > 0$ and H > 0. This has direct implications for the tail behavior of other well known branching recursions.

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1. Introduction

In recent years considerable attention [1, 2, 3, 4, 5, 6, 7, 8] has been given to the characterization and analysis of the solutions to the non homogeneous

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linear equation

$$R_L \stackrel{\mathcal{D}}{=} \sum_{i=1}^N C_i R_{L,i} + Q, \qquad (1)$$

where $(Q, N, C_1, C_2, ...)$ is a real-valued random vector with $N \in \mathbb{N} \cup \{\infty\}$, P(|Q| > 0) > 0, and $\{R_{L,i}\}_{i \in \mathbb{N}}$ is a sequence of i.i.d. random variables independent of $(Q, N, C_1, C_2, ...)$ having the same distribution as R_L . Equation (1) has applications in a wide variety of fields, including the analysis of divide and conquer algorithms [9, 10], e.g. Quicksort [11]; the analysis of the PageRank algorithm [12, 1]; and kinetic gas theory [8]. Our work in [2, 3] shows that the so-called endogenous solution, as termed in [13], of (1), under the natural main root condition $E\left[\sum_{i=1}^{N} |C_i|^{\alpha}\right] = 1$ with positive derivative $0 < E\left[\sum_{i=1}^{N} |C_i|^{\alpha} \log |C_i|\right] < \infty$ for some $\alpha > 0$, has the power tail behavior,

$$P(|R_L| > t) \sim H_L t^{-\alpha}, \qquad t \to \infty,$$

where $0 \leq H_L < \infty$. The main tool used in deriving this result was a generalization of Goldie's Implicit Renewal Theorem [14] to weighted branching trees.

Motivated by a different set of applications, we study in this paper the maximum recursion on trees given by

$$R \stackrel{\mathcal{D}}{=} \left(\bigvee_{i=1}^{N} C_i R_i\right) \lor Q,\tag{2}$$

where $(Q, N, C_1, C_2, ...)$ is a random vector with $N \in \mathbb{N} \cup \{\infty\}$, nonnegative weights $\{C_i\}$, and P(|Q| > 0) > 0, and $\{R_i\}_{i \in \mathbb{N}}$ is a sequence of i.i.d. random variables independent of $(Q, N, C_1, C_2, ...)$ having the same distribution as R. Here and throughout the paper we use $x \vee y$ and $x \wedge y$ to denote the maximum and minimum, respectively, of x and y. We point out that by taking the logarithm in (2) when Q > 0 a.s., we obtain the additive equivalent

$$X \stackrel{\mathcal{D}}{=} \bigvee_{i=1}^{N} (Y_i + X_i) \lor V, \tag{3}$$

where $X = \log R$, $Y_i = \log C_i$, $V = \log Q$, and the $\{X_i\}_{i \in \mathbb{N}}$ are i.i.d. copies of X, independent of $(V, N, Y_1, Y_2, ...)$. Note that for $N \equiv 1$ and $V \equiv 0$, (3) reduces to the classical Lindley's equation, satisfied by the reflected random walk; and when $V \not\equiv 0$, the recursion corresponds to a random walk reflected on a random barrier. In general, the preceding additive equation has been studied in the literature of branching random walks (see $[13], \S4.2$). Recursion (3) was termed "high-order Lindley equation" and studied in the context of queues with synchronization in [15]. Unlike the classical Lindley equation, it was shown in [15] that (3) can have multiple solutions. A more complete analysis of the existence and the characterization of the entire family of solutions was carried out in [16] (e.g., see Theorem 1 in [16]). In addition, it can be shown that the study of (3) arises in the context of today's massively parallel computing, e.g., consider a job that is split into smaller pieces which are sent randomly to different processors, and these pieces need to be synchronized in order to complete their processing. In addition to these applications, a better understanding of (2) immediately leads to important insights to other max-plus branching recursions. More precisely, for the case of nonnegative weights, (2) is a natural lower bound for many other recursions on trees [13].

For all of the reasons described above, we study in this paper the tail behavior of the minimal/endogenous solution to the maximum recursion in (2) (or (3)). Furthermore, we point out that under iterations of the fixedpoint equation (2) (or (3)), the minimal/endogenous solution is the primary limiting value, unless one starts with very specific initial distributions (see Theorem 1(ii) in [16]); we will discuss this in more detail in Section 3.

Our first main result, stated in Theorem 3.4, describes the tail behavior of the minimal/endogenous solution to the maximum recursion (2) (or (3)). In this regard, the application of the Implicit Renewal Theorem on Trees (see Theorem 3.4 [3]), under the natural conditions $E\left[\sum_{i=1}^{N} C_{i}^{\alpha}\right] = 1$ and $0 < E\left[\sum_{i=1}^{N} C_{i}^{\alpha} \log C_{i}\right] < \infty$ for some $\alpha > 0$, readily gives $P(R > t) \sim Ht^{-\alpha}, \qquad t \to \infty,$ (4)

where
$$0 \leq H < \infty$$
. However, the main difficulty in establishing the power-
law behavior lies in proving that $H > 0$. Unlike in the linear case, it is not
clear that this constant should be positive at all, since at first glance the

$$E\left[(Q^{+})^{\alpha} \vee \bigvee_{i=1}^{N} (C_{i}R_{i}^{+})^{\alpha} - \sum_{i=1}^{N} (C_{i}R_{i}^{+})^{\alpha}\right],$$

expression which determines H in Theorem 3.4 of Section 3,

appears just as likely to be negative. Hence, our first main contribution lies in a new sample-path construction showing that H > 0 under no additional assumptions (besides those needed for the application of Theorem 3.4 in [3]). Observe that in the additive case of equation (3), our result yields the exponential asymptotics $P(X > y) \sim He^{-\alpha y}$, which is the generalization of the well known Cramér-Lundberg approximation. The latter is widely used in insurance risk theory and queueing.

Furthermore, as an immediate corollary one obtains the strict positivity of H_L in the linear case with nonnegative $(Q, N, C_1, C_2, ...)$. In this setting, the work in [2] used a straightforward convexity argument to show that $H_L > 0$ for $\alpha \ge 1$, but the corresponding question for $\alpha \in (0, 1)$ was left open. The strict positivity of H_L for $\alpha \in (0, 1)$ was recently resolved in [7] as part of the more general real-valued case, but under additional assumptions that include $E\left[\sum_{i=1}^N C_i^{\alpha+\epsilon}\right] < \infty$. Note that in the additive equation (3), this extra moment assumption corresponds to the finiteness of $\alpha + \epsilon$ exponential moments of the $\{Y_i\}$. Since the new results on the maximum hold without such additional assumptions, Theorem 3.4 fully completes the prior work for nonnegative (Q, N, C_1, C_2, \ldots) .

We now go back to the linear recursion (1) with real-valued weights $(Q, C_1, C_2, ...)$, which has recently been considered in [5, 3, 7] (see also [8] for the multivariate case). The work in [3] establishes the Implicit Renewal Theorem on Trees for the real-valued case and shows that, under the usual conditions $E\left[\sum_{i=1}^{N} |C_i|^{\alpha}\right] = 1$ and $0 < E\left[\sum_{i=1}^{N} |C_i|^{\alpha} \log |C_i|\right] < \infty$, the endogenous solution to (1) has a power tail behavior of the form $H_L t^{-\alpha}$, $H_L \geq 0$. In that paper the strict positivity of H_L in its full generality remained open. It was this open problem that motivated the work in [7], where it was shown, using complex analysis and analytical functions, that $H_L > 0$ under the additional assumptions $N < \infty$ a.s., $E\left[\sum_{i=1}^{N} |C_i|^{\alpha+\epsilon}\right] < \infty$ and $E\left[\left(\sum_{i=1}^{N} |C_i|\right)^{\alpha+\epsilon}\right] < \infty$.

In this paper, we revisit the problem of the strict positivity of H_L for the general real-valued case using our result on the maximum equation (2) (with nonnegative weights $\{C_i\}$), under no additional assumptions on the vector $(N, C_1, C_2, ...)$ besides those needed for Theorem 3.4 in [3]. However, we do require that Q does not reduce to a constant given $(N, C_1, C_2, ...)$, i.e., Q can be a random, but not *deterministic*, function of $(N, C_1, C_2, ...)$. Our

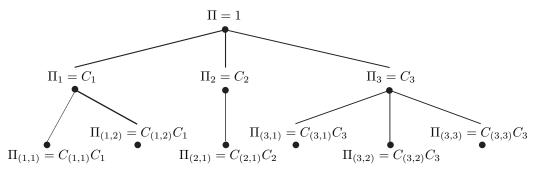


Figure 1: Weighted branching tree

main set of arguments is based on Lévy's symmetrization approach.

The paper is organized as follows. Section 2 includes a brief description of the weighted branching process. Section 3 contains our first main result about the asymptotic behavior of the minimal/endogenous solution to the maximum recursion (2), including the strict positivity of H. Section 4 presents our proof of the positivity of the constant H_L for the general mixed-sign linear recursion (1).

2. Model description

We use the model from [3] for defining a weighted branching tree. To this end, let $\mathbb{N}_+ = \{1, 2, 3, ...\}$ be the set of positive integers and let $U = \bigcup_{k=0}^{\infty} (\mathbb{N}_+)^k$ be the set of all finite sequences $\mathbf{i} = (i_1, i_2, ..., i_n) \in U$, where by convention $\mathbb{N}^0_+ = \{\emptyset\}$ contains the null sequence \emptyset . To ease the exposition, for a sequence $\mathbf{i} = (i_1, i_2, ..., i_k) \in U$ we write $\mathbf{i} | n = (i_1, i_2, ..., i_n)$, provided $k \ge n$, and $\mathbf{i} | 0 = \emptyset$ to denote the index truncation at level $n, n \ge 0$. Also, for $\mathbf{i} \in A_1$ we simply use the notation $\mathbf{i} = i_1$, that is, without the parenthesis. Similarly, for $\mathbf{i} = (i_1, ..., i_n)$ we will use $(\mathbf{i}, j) = (i_1, ..., i_n, j)$ to denote the index concatenation operation, if $\mathbf{i} = \emptyset$, then $(\mathbf{i}, j) = j$.

To iteratively construct the weighted branching tree \mathcal{T} , let $\{(N_{\mathbf{i}}, C_{(\mathbf{i},1)}, C_{(\mathbf{i},2)}, \dots)\}_{\mathbf{i}\in U}$ be a sequence of i.i.d. random vectors. The random variables $\{N_{\mathbf{i}}\}_{\mathbf{i}\in U}$ in this sequence define the structure of the tree as follows; set $N = N_{\emptyset}$. Let $A_0 = \{\emptyset\}$,

$$A_{1} = \{ i \in \mathbb{N} : 1 \le i \le N \}, \text{ and} A_{n} = \{ (\mathbf{i}, i_{n}) \in U : \mathbf{i} \in A_{n-1}, 1 \le i_{n} \le N_{\mathbf{i}} \}, n \ge 2,$$
(5)

be the set of individuals in the *n*th generation. Next, assign to each node **i** in the tree a weight $\Pi_{\mathbf{i}}$ according to the recursion

$$\Pi_{i_1} = C_{i_1}, \qquad \Pi_{(i_1,\dots,i_n)} = C_{(i_1,\dots,i_n)} \Pi_{(i_1,\dots,i_{n-1})}, \qquad n \ge 2,$$

where $\Pi = \Pi_{\emptyset} \equiv 1$ is the weight of the root node. See Figure 1.

3. The maximum recursion: $R = \left(\bigvee_{i=1}^N C_i R_i \right) \lor Q$

In this section, we study the maximum fixed-point equation given by

$$R \stackrel{\mathcal{D}}{=} \left(\bigvee_{i=1}^{N} C_i R_i\right) \lor Q,\tag{6}$$

where $(Q, N, C_1, C_2, ...)$ is a random vector with $N \in \mathbb{N} \cup \{\infty\}, \{C_i\} \ge 0$ and P(|Q| > 0) > 0, and $\{R_i\}_{i \in \mathbb{N}}$ is a sequence of i.i.d. random variables independent of $(Q, N, C_1, C_2, ...)$ having the same distribution as R. As already mentioned, the additive version of (6), given in (3), was termed "high-order Lindley equation" and studied in the context of queues with synchronization in [15]. The full characterization of its multiple solutions was given in [16]. More recently, a related recursion where $Q \equiv 0, N = \infty$, and the $\{C_i\}$ are real valued deterministic constants, has been analyzed in [17]. The more closely related case of $Q \equiv 0$ and $\{C_i\} \ge 0$ being random was studied earlier in [18]. For this and other max-plus equations appearing in a variety of applications see the survey by [13].

Using standard arguments, we start by constructing an endogenous solution to (6) on a tree and then we show that this solution is finite a.s. and represents the unique limit under iterations provided that the initial values and the weights satisfy appropriate moment conditions.

Following the notation of Section 2, define the process

$$V_n = \bigvee_{\mathbf{i} \in A_n} Q_{\mathbf{i}} \Pi_{\mathbf{i}}, \qquad n \ge 0, \tag{7}$$

on the weighted branching tree \mathcal{T} . The convention throughout the paper is that $(Q, N, C_1, C_2, ...) = (Q_{\emptyset}, N_{\emptyset}, C_{(\emptyset,1)}, C_{(\emptyset,2)}, ...)$ denotes the random vector corresponding to the root node. Next, define the process $\{R^{(n)}\}_{n\geq 0}$ according to

$$R^{(n)} = \bigvee_{k=0}^{n} V_k, \qquad n \ge 0.$$

It is not hard to see that $R^{(n)}$ satisfies the recursion

$$R^{(n)} = \left(\bigvee_{j=1}^{N_{\emptyset}} C_{(\emptyset,j)} R_j^{(n-1)}\right) \lor Q_{\emptyset} = \left(\bigvee_{j=1}^{N} C_j R_j^{(n-1)}\right) \lor Q,$$
(8)

where $\{R_j^{(n-1)}\}\$ are independent copies of $R^{(n-1)}$ corresponding to the tree starting with individual j in the first generation and ending on the *n*th generation. One can also verify that

$$V_n = \bigvee_{k=1}^{N_{\emptyset}} C_{(\emptyset,k)} \bigvee_{(k,i_2,\dots,i_n) \in A_n} Q_{(k,i_2,\dots,i_n)} \prod_{j=2}^n C_{(k,i_2,\dots,i_j)} \stackrel{\mathcal{D}}{=} \bigvee_{k=1}^N C_k V_{(n-1),k},$$

where $\{V_{(n-1),k}\}$ is a sequence of i.i.d. random variables independent of $(N, C_1, C_2, ...)$ and having the same distribution as V_{n-1} .

We now define the random variable R according to

$$R \triangleq \lim_{n \to \infty} R^{(n)} = \bigvee_{k=0}^{\infty} V_k.$$
(9)

Note that $R^{(n)}$ is monotone increasing sample-pathwise, so R is well defined. Also, by monotonicity of $R^{(n)}$ and (8), we obtain that R solves

$$R = \left(\bigvee_{j=1}^{N_{\emptyset}} C_{(\emptyset,j)} R_j^{(\infty)}\right) \lor Q_{\emptyset} = \left(\bigvee_{j=1}^{N} C_j R_j^{(\infty)}\right) \lor Q,$$

where $\{R_j^{(\infty)}\}_{j\in\mathbb{N}}$ are i.i.d. copies of R, independent of (Q, N, C_1, C_2, \ldots) , see also Section 2 in [16]. Clearly this implies that R, as defined by (9), is a solution in distribution to (6). However, this solution might be ∞ . Next, we establish in the following lemma the finiteness of moments of R, and in particular that $R < \infty$ a.s.; its proof uses standard contraction arguments but is included for completeness; e.g. see Theorem 6 (i) in [16]. Conditions under which R is infinite a.s. can be found in Corollary 4 in [16].

Lemma 3.1. Assume that $\rho_{\beta} = E\left[\sum_{i=1}^{N} C_{i}^{\beta}\right] < 1$ and $E[|Q|^{\beta}] < \infty$ for some $\beta > 0$. Then, $E[|R|^{\gamma}] < \infty$ for all $0 < \gamma \leq \beta$, and in particular, $|R| < \infty$ a.s. Moreover, if $\beta \geq 1$, $R^{(n)} \xrightarrow{L_{\beta}} R$, where L_{β} stands for convergence in $(E|\cdot|^{\beta})^{1/\beta}$ norm.

Proof. It follows immediately that

$$E\left[|R|^{\beta}\right] \le E\left[\sum_{k=0}^{\infty} \sum_{\mathbf{i} \in A_k} |Q_{\mathbf{i}}|^{\beta} \Pi_{\mathbf{i}}^{\beta}\right] = \sum_{k=0}^{\infty} E[|Q|^{\beta}] \rho_{\beta}^k = \frac{E[|Q|^{\beta}]}{1 - \rho_{\beta}} < \infty,$$

That $R^{(n)} \xrightarrow{L_{\beta}} R$ whenever $\beta \geq 1$ follows from noting that $|R^{(n)} - R|^{\beta} \leq |\bigvee_{k=n+1}^{\infty} V_k|^{\beta}$ and the same arguments used above. \Box

Although this paper focuses only on the solution R defined by (9), it is important to mention that equation (6) can have multiple solutions, as the work in [16] describes. The solution R receives the name "endogenous" since it is constructed explicitly from the weighted branching tree, and the name "minimal" since it is the stochastically smallest solution, in the sense that any other solution R' to (6) satisfies $P(R' > t) \leq P(R > t)$ for all t > 0. For the case when $Q \geq 0$ and there exists a unique v > 0 such that $E\left[\sum_{i=1}^{N} C_i^v\right] = 1$ and $-\infty < E\left[\sum_{i=1}^{N} C_i^v \log C_i\right] < 0$ (referred to as the "regular case"), Theorem 1 (ii) and (iii) of [16] characterizes the entire family of solutions to (6). Moreover, under some additional technical conditions, all other solutions to (6) are given in terms of R ($M = \log R$ in [16]) and the limit W(v) of the martingale $W_k(v) = \sum_{i \in A_k} \prod_i^v$. To better understand the nature of these other solutions, as well as to highlight the importance of the endogenous/minimal solution R, we will next define the process $\{R_n^*\}$ that is obtained from iterating equation (6) starting from an initial value R_0^* .

Let

$$R_n^* \triangleq R^{(n-1)} \lor V_n(R_0^*), \qquad n \ge 1,$$

where

$$V_n(R_0^*) = \bigvee_{\mathbf{i} \in A_n} R_{0,\mathbf{i}}^* \Pi_{\mathbf{i}},\tag{10}$$

and $\{R_{0,i}^*\}_{i \in U}$ are i.i.d. copies of an initial value R_0^* , independent of the entire weighted tree $\mathcal{T}_{Q,C}$. R_0^* is referred to as the "terminal" value in [16] $(T = \log R_0^*, R_0^* \ge 0)$ since it corresponds to the value of the leaves in the weighted branching tree with finitely many generations. It follows from (8) and (10) that

$$R_{n+1}^{*} = \bigvee_{j=1}^{N} C_{j} \left(R_{j}^{(n-1)} \vee \bigvee_{\mathbf{i} \in A_{n,j}} R_{0,\mathbf{i}}^{*} \prod_{k=2}^{n} C_{(j,\dots,i_{k})} \right) \vee Q = \bigvee_{j=1}^{N} C_{j} R_{n,j}^{*} \vee Q,$$

where $\{R_j^{(n-1)}\}\$ are independent copies of $R^{(n-1)}$ corresponding to the tree starting with individual j in the first generation and ending on the nth generation, and $A_{n,j}$ is the set of all nodes in the (n + 1)th generation that are descendants of individual j in the first generation. Moreover, $\{R_{n,j}^*\}\$ are i.i.d. copies of R_n^* , and thus, R_n^* is equal in distribution to the process obtained by iterating (6) with an initial condition R_0^* . This process can be shown to converge in distribution to R for any initial condition R_0^* satisfying the following moment condition (see also Theorem 9 in [16]).

Lemma 3.2. Suppose $E[|Q|^{\beta}], E[|R_0^*|^{\beta}] < \infty$ and $\rho_{\beta} < 1$ for some $\beta > 0$, then

$$R_n^* \Rightarrow R_n$$

with $E[|R|^{\beta}] < \infty$. Furthermore, under these assumptions, the distribution of R is the unique solution with finite β -moment to recursion (6).

Proof. The result is immediate from Lemma 5.2 in [2] by noting that

$$|V_n(R_0^*)| \le \bigvee_{\mathbf{i}\in A_n} \left| R_{0,\mathbf{i}}^* \Pi_{\mathbf{i}} \right|.$$

Remarks 3.3. (a) Lemma 3.2 establishes a certain type of uniqueness of the solution to (6), in the sense that R is the only possible limit for the iterative process $\{R_n^*\}$ for any initial value R_0^* possessing finite β moment with $\rho_\beta < 1$. It is therefore to be expected that all other solutions to the maximum recursion must arise from violating this assumption. (b) Theorem 1 (ii) of [16] states that in the regular case (see the comments after Lemma 3.1), if $R_0^* \ge 0$ and $\lim_{t\to\infty} t^{\upsilon} P(R_0^* > t) = \gamma$ ($\upsilon < \alpha$), then $R_n^* \Rightarrow R(\gamma)$, where

$$P(R(\gamma) \le t) = E\left[1(R \le t)e^{-\gamma W(\upsilon)t^{\upsilon}}\right].$$

Moreover, $R(\gamma)$ solves (6) provided $R < \infty$ a.s. and $E\left[W_1(\upsilon)\log^+ W_1(\upsilon)\right] < \infty$.

Now we are ready to state the main result of this section, which characterizes the asymptotic behavior of R.

Theorem 3.4. Let $(Q, N, C_1, C_2, ...)$ be a random vector with $N \in \mathbb{N} \cup \{\infty\}$, $\{C_i\} \ge 0$ and P(|Q| > 0) > 0, and R be the solution to (6) given by (9).

Suppose that there exists $j \ge 1$ with $P(N \ge j, C_j > 0) > 0$ such that the measure $P(\log C_j \in du, C_j > 0, N \ge j)$ is nonarithmetic, and that for some $\alpha > 0$, $E[|Q|^{\alpha}] < \infty$, $0 < E\left[\sum_{i=1}^{N} C_i^{\alpha} \log C_i\right] < \infty$ and $E\left[\sum_{i=1}^{N} C_i^{\alpha}\right] = 1$. In addition, assume

1.
$$E\left[\left(\sum_{i=1}^{N} C_{i}\right)^{\alpha}\right] < \infty$$
, if $\alpha > 1$; or,
2. $E\left[\left(\sum_{i=1}^{N} C_{i}^{\alpha/(1+\epsilon)}\right)^{1+\epsilon}\right] < \infty$ for some $0 < \epsilon < 1$, if $0 < \alpha \le 1$.

Then,

$$P(R > t) \sim Ht^{-\alpha}, \qquad P(R < -t) = o(t^{-\alpha}), \qquad t \to \infty$$

where $0 \leq H < \infty$ is given by

$$H = \frac{1}{E\left[\sum_{i=1}^{N} C_i^{\alpha} \log C_i\right]} \int_0^{\infty} v^{\alpha-1} \left(P(R > v) - E\left[\sum_{i=1}^{N} 1(C_i R_i > v)\right] \right) dv$$
$$= \frac{E\left[(Q^+)^{\alpha} \lor \bigvee_{i=1}^{N} (C_i R_i^+)^{\alpha} - \sum_{i=1}^{N} (C_i R_i^+)^{\alpha}\right]}{\alpha E\left[\sum_{i=1}^{N} C_i^{\alpha} \log C_i\right]}.$$

Furthermore, H > 0 if and only if $P(Q^+ > 0) > 0$.

Remarks 3.5. (a) The condition $E[|Q|^{\alpha}] < \infty$ is only needed to obtain the result about the negative tail. The result about the positive tail P(R > t) only requires $E[(Q^+)^{\alpha}] < \infty$. (b) The equivalent result for the lattice case can be obtained by using the corresponding Implicit Renewal Theorem on Trees in [3]. (c) Corollary 5 in [16] provides upper bounds for the tail behavior of any finite solution to the maximum equation (6).

Proof. The first part of the proof about the right tail, P(R > t), will follow from an application of the Implicit Renewal Theorem on Trees, Theorem 3.4 in [3], once we verify the finiteness of

$$\int_0^\infty \left| P(R>t) - E\left[\sum_{i=1}^N 1(C_i R_i > t)\right] \right| t^{\alpha - 1} dt.$$
(11)

To see that (11) is indeed finite, note that by Lemma 4.10 in [3] we have that

$$0 \le \int_0^\infty \left(E\left[\sum_{i=1}^N 1(C_i R_i > t)\right] - P\left(\bigvee_{i=1}^N C_i R_i > t\right) \right) t^{\alpha - 1} dt < \infty.$$

Also, since $R^* \triangleq \left(\bigvee_{i=1}^N C_i R_i\right) \lor Q \ge \bigvee_{i=1}^N C_i R_i$, then

$$0 \leq \int_0^\infty \left(P(R^* > t) - P\left(\bigvee_{i=1}^N C_i R_i > t\right) \right) t^{\alpha - 1} dt$$

$$= \frac{1}{\alpha} E\left[\left(\left(\left(\bigvee_{i=1}^N C_i R_i\right) \lor Q\right)^+ \right)^\alpha - \left(\left(\bigvee_{i=1}^N C_i R_i\right)^+ \right)^\alpha \right]$$

$$= \frac{1}{\alpha} E\left[(Q^+)^\alpha \lor \bigvee_{i=1}^N (C_i R_i^+)^\alpha - \bigvee_{i=1}^N (C_i R_i^+)^\alpha \right]$$

$$\leq \frac{1}{\alpha} E[(Q^+)^\alpha].$$

Combining these two observations gives that (11) is finite, and by Theorem 3.4 (a) in [3] we obtain the result with the integral representation of H. To derive the second expression for H follow the same steps used at the end of the proof of Theorem 4.1 in [2].

For the negative tail, P(R < -t), simply note that

$$P(R < -t) = P\left(\left(\bigvee_{i=1}^{N} C_{i}R_{i}\right) \lor Q < -t\right) \le P(Q < -t)$$
$$\le P(|Q| > t) \le E\left[|Q|^{\alpha}1(|Q|^{\alpha} > t)\right]t^{-\alpha},$$

where in the last step we used Markov's inequality. Since $E[|Q|^{\alpha}] < \infty$, then $E[|Q|^{\alpha}1(|Q|^{\alpha} > t)] = o(1)$ as $t \to \infty$, proving the result.

The rest of the proof is devoted to showing that the constant H > 0 if and only if $P(Q^+ > 0) > 0$. Note that if $Q \le 0$ a.s. then $R^+ = 0$ and therefore H = 0, so it only remains to show that H > 0 whenever $P(Q^+ > 0) > 0$. Hence, assume from now on that $P(Q^+ > 0) > 0$.

The main idea of the proof is to construct a minorizing random variable for R for which we can directly estimate the expectation appearing in the numerator of H. We start by fixing $0 < \delta < E[(Q^+)^{\alpha}] \wedge 1$ and choosing $\alpha/2 < \beta < \alpha$ and q > 0 such that $\rho_{\beta} < 1$, $E\left[\left(\sum_{i=1}^{N} C_{i}^{\beta}\right)^{\alpha/\beta}\right] < \infty$, and $E[(Q^+)^{\alpha}1(Q^+ > q)] < \delta/6$; define $K = \beta^{-1} \int_{0}^{\infty} (e^{-u} - 1 + u) u^{-\alpha/\beta - 1} du < \infty$. Note that such β always exists under the assumptions of the theorem, since when $0 < \alpha \leq 1$ we have that for any $\alpha/(1+\epsilon) \leq \beta < \alpha$,

$$E\left[\left(\sum_{i=1}^{N} C_{i}^{\beta}\right)^{\alpha/\beta}\right] \leq E\left[\left(\sum_{i=1}^{N} C_{i}^{\frac{\alpha}{1+\epsilon}}\right)^{\frac{\alpha}{\beta} \cdot \frac{(1+\epsilon)\beta}{\alpha}}\right] = E\left[\left(\sum_{i=1}^{N} C_{i}^{\frac{\alpha}{1+\epsilon}}\right)^{1+\epsilon}\right] < \infty,$$

and when $\alpha > 1$ we have that for any $1 \leq \beta < \alpha$,

$$E\left[\left(\sum_{i=1}^{N} C_{i}^{\beta}\right)^{\alpha/\beta}\right] \leq E\left[\left(\sum_{i=1}^{N} C_{i}\right)^{\frac{\alpha}{\beta}\cdot\beta}\right] = E\left[\left(\sum_{i=1}^{N} C_{i}\right)^{\alpha}\right] < \infty.$$

Now let $\{D_{\mathbf{i},j} : \mathbf{i} \in U, 1 \leq j \leq r\}$ be nonnegative i.i.d. random variables, independent of \mathcal{T} , having the same distribution as D, where D satisfies

$$0 \le D \le d$$
 a.s., $E[D^{\alpha}] = 1$ and $E[D^{\beta}] < 1$,

(e.g., take D to have density $f(x) = (\alpha/2)x^{\alpha-1}1(0 \le x \le 2^{1/\alpha}))$. For each $\mathbf{i} \in \mathcal{T}$ define the random variable

$$\mathcal{Q}_{\mathbf{i}} = Q_{\mathbf{i}} \prod_{j=1}^{r} D_{(\mathbf{i},j)},$$

where $r \in \mathbb{N}$ is such that

$$\frac{q^{\alpha} E[(Q^+)^{\beta}]}{(\delta/6)(1-\rho_{\beta})} (E[D^{\beta}])^r < \delta/6$$

and

$$KE\left[\left(\sum_{i=1}^{N} C_{i}^{\beta}\right)^{\alpha/\beta}\right] \left(\frac{E[(Q^{+})^{\beta}]}{1-\rho_{\beta}}\right)^{\alpha/\beta} (E[D^{\beta}])^{\alpha r/\beta} < \alpha\delta/2.$$

Let

$$\mathcal{R} = \bigvee_{\mathbf{i} \in \mathcal{T}} \mathcal{Q}_{\mathbf{i}} \Pi_{\mathbf{i}},$$

and note that for any t > 0,

$$P(R > t) = P(R^{+} > t) = P\left(\left(\bigvee_{\mathbf{i}\in\mathcal{T}}Q_{\mathbf{i}}\Pi_{\mathbf{i}}\right)^{+} > t\right) = P\left(\bigvee_{\mathbf{i}\in\mathcal{T}}d^{r}Q_{\mathbf{i}}^{+}\Pi_{\mathbf{i}} > d^{r}t\right)$$
$$\geq P\left(\bigvee_{\mathbf{i}\in\mathcal{T}}\prod_{j=1}^{r}D_{(\mathbf{i},j)}Q_{\mathbf{i}}^{+}\Pi_{\mathbf{i}} > d^{r}t\right) = P\left(\bigvee_{\mathbf{i}\in\mathcal{T}}Q_{\mathbf{i}}^{+}\Pi_{\mathbf{i}} > d^{r}t\right)$$
$$= P(\mathcal{R} > d^{r}t).$$

We now apply the first part of this theorem to the new random variable \mathcal{R} to obtain

$$P(\mathcal{R} > v) \sim \frac{E\left[(\mathcal{Q}^+)^{\alpha} \vee \bigvee_{i=1}^N (C_i \mathcal{R}_i^+)^{\alpha} - \sum_{i=1}^N (C_i \mathcal{R}_i^+)^{\alpha}\right]}{\alpha E\left[\sum_{i=1}^N C_i^{\alpha} \log C_i\right]} \cdot v^{-\alpha}$$

as $v \to \infty$. The positivity of H will then follow once we show

$$\mathcal{E} \triangleq E\left[(\mathcal{Q}^+)^{\alpha} \lor \bigvee_{i=1}^N (C_i \mathcal{R}_i^+)^{\alpha} - \sum_{i=1}^N (C_i \mathcal{R}_i^+)^{\alpha} \right] > 0.$$

We start by writing \mathcal{E} as

$$\mathcal{E} = E\left[\left((\mathcal{Q}^+)^{\alpha} - \bigvee_{i=1}^N (C_i \mathcal{R}_i^+)^{\alpha}\right)^+\right] - E\left[\sum_{i=1}^N (C_i \mathcal{R}_i^+)^{\alpha} - \bigvee_{i=1}^N (C_i \mathcal{R}_i^+)^{\alpha}\right]$$
$$\triangleq \mathcal{E}_1 - \mathcal{E}_2.$$

To analyze \mathcal{E}_1 note that

$$\mathcal{E}_{1} \geq E\left[\left((\mathcal{Q}^{+})^{\alpha}1(Q^{+} \leq q) - \bigvee_{i=1}^{N}(C_{i}\mathcal{R}_{i}^{+})^{\alpha}\right)^{+}1\left(\bigvee_{i=1}^{N}(C_{i}\mathcal{R}_{i}^{+})^{\alpha} \leq \delta/6\right)\right]$$

$$\geq E\left[\left((\mathcal{Q}^{+})^{\alpha}1(Q^{+} \leq q) - \delta/6\right)^{+}1\left(\bigvee_{i=1}^{N}(C_{i}\mathcal{R}_{i}^{+})^{\alpha} > \delta/6\right)\right]$$

$$\geq E\left[\left((\mathcal{Q}^{+})^{\alpha}1(Q^{+} \leq q)\right] - \frac{\delta}{6} - q^{\alpha}E\left[\left(\prod_{j=1}^{r}D_{j}^{\alpha}\right)1\left(\bigvee_{i=1}^{N}(C_{i}\mathcal{R}_{i}^{+})^{\alpha} > \delta/6\right)\right]$$

$$\equiv E\left[\left((\mathcal{Q}^{+})^{\alpha}1(Q^{+} \leq q)\right] - \frac{\delta}{6} - q^{\alpha}E\left[\left(\prod_{j=1}^{r}D_{j}^{\alpha}\right)1\left(\bigvee_{i=1}^{N}(C_{i}\mathcal{R}_{i}^{+})^{\alpha} > \delta/6\right)\right]$$

 $= E\left[(Q^{+})^{\alpha}\right] - E\left[(Q^{+})^{\alpha}1(Q^{+} > q)\right] - \frac{\delta}{6} - q^{\alpha}P\left(\bigvee_{i=1}^{n} (C_{i}\mathcal{R}_{i}^{+})^{\alpha} > \delta/6\right),$

where in the last equality we used the observation that $(\mathcal{Q}^+)^{\alpha} = (Q^+)^{\alpha} \prod_{j=1}^r D_j^{\alpha}$,

where $\prod_{j=1}^{r} D_{j}^{\alpha}$ is independent of \mathcal{T} , and $E\left[\prod_{j=1}^{r} D_{j}^{\alpha}\right] = 1$. It follows that

$$\mathcal{E}_{1} \geq E[(Q^{+})^{\alpha}] - \frac{\delta}{3} - q^{\alpha}P\left(\bigvee_{i=1}^{N} (C_{i}\mathcal{R}_{i}^{+})^{\beta} > (\delta/6)^{\beta/\alpha}\right)$$
$$\geq E[(Q^{+})^{\alpha}] - \frac{\delta}{3} - \frac{q^{\alpha}}{(\delta/6)^{\beta/\alpha}}E\left[\sum_{i=1}^{N} (C_{i}\mathcal{R}_{i}^{+})^{\beta}\right] \qquad \text{(by Markov's inequality)}$$
$$= E[(Q^{+})^{\alpha}] - \frac{\delta}{3} - \frac{q^{\alpha}\rho_{\beta}}{(\delta/6)^{\beta/\alpha}}E\left[(\mathcal{R}^{+})^{\beta}\right].$$

By the same arguments used in the proof of Lemma 3.1,

$$E[(\mathcal{R}^+)^\beta] \le E\left[\sum_{k=0}^\infty \sum_{\mathbf{i} \in A_k} (\mathcal{Q}_{\mathbf{i}}^+)^\beta \Pi_{\mathbf{i}}^\beta\right] = \frac{E\left[(\mathcal{Q}^+)^\beta\right]}{1-\rho_\beta} = \frac{\left(E[D^\beta]\right)^r E[(Q^+)^\beta]}{1-\rho_\beta}.$$
(12)

Our choice of r now guarantees that

$$\mathcal{E}_1 \ge E[(Q^+)^{\alpha}] - \frac{\delta}{3} - \frac{q^{\alpha} E[(Q^+)^{\beta}]}{(\delta/6)(1-\rho_{\beta})} (E[D^{\beta}])^r > E[(Q^+)^{\alpha}] - \frac{\delta}{2}$$

It remains to bound \mathcal{E}_2 . Follow the same steps as in the proof of Lemma 4.6 in [2] to obtain

$$\alpha \mathcal{E}_{2} = \int_{0}^{\infty} E\left[\sum_{i=1}^{N} 1(C_{i}\mathcal{R}_{i}^{+} > t) - 1\left(\bigvee_{i=1}^{N} C_{i}\mathcal{R}_{i}^{+} > t\right)\right] t^{\alpha-1} dt$$

$$\leq E\left[\beta^{-1}\left(E[(\mathcal{R}^{+})^{\beta}]\sum_{i=1}^{N} C_{i}^{\beta}\right)^{\alpha/\beta} \int_{0}^{\infty} \left(e^{-u} - 1 + u\right) u^{-\alpha/\beta-1} du\right]$$

$$\leq KE\left[\left(\sum_{i=1}^{N} C_{i}^{\beta}\right)^{\alpha/\beta}\right]\left(\frac{\left(E[D^{\beta}]\right)^{r} E[(Q^{+})^{\beta}]}{1 - \rho_{\beta}}\right)^{\alpha/\beta} \quad (by (12)).$$

Our choice of r now gives $\mathcal{E}_2 < \delta/2$. We conclude that

$$\mathcal{E} > E[(Q^+)^{\alpha}] - \delta > 0.$$

4. The linear recursion

In this section of the paper we explain how Theorem 3.4, which establishes the power-law behavior of the endogenous solution R to the maximum equation (6), can be used to show that the constant H_L given by Theorem 4.6 in [3] is strictly positive.

Consider the linear equation

$$R_L \stackrel{\mathcal{D}}{=} \sum_{i=1}^{N} C_i R_{L,i} + Q, \qquad (13)$$

where $(Q, N, C_1, C_2, ...)$ is a real-valued random vector with $N \in \mathbb{N} \cup \{\infty\}$ and P(|Q| > 0) > 0, and $\{R_{L,i}\}_{i \in \mathbb{N}}$ is a sequence of i.i.d. random variables independent of $(Q, N, C_1, C_2, ...)$ having the same distribution as R_L . The results in this section refer to the endogenous solution given by

$$R_L = \sum_{k=0}^{\infty} \sum_{\mathbf{i} \in A_k} \Pi_{\mathbf{i}} Q_{\mathbf{i}}.$$
 (14)

We refer the reader to Section 4 in [3] for detailed conditions under which R_L is well defined and how it solves (13). We point out that all our results for the maximum recursion assume that the weights $\{C_i\}$ are nonnegative, so the connection between R_L and R is much more difficult to make in this case.

As mentioned in the introduction, the idea behind our proof lies in first considering what we call a "symmetric tree", and using a novel argument to show that the corresponding endogenous solution to the linear recursion on this symmetric tree follows a power-law asymptotic behavior with a strictly positive constant of proportionality. The second step is to construct a symmetric tree using two coupled versions of a general (non-symmetric) tree and show how the solution to the general case is lower bounded by the symmetric solution. The first of these two steps is given in the following proposition.

Proposition 4.1. Let $(Q, N, C_1, C_2, ...)$ be a random vector with $N \in \mathbb{N} \cup \{\infty\}$ and P(|Q| > 0) > 0, and R_L be the solution to (13) given by (14). Assume that for some $0 < \beta \le 1$, $E[|Q|^{\beta}] < \infty$ and $E\left[\sum_{j=1}^{N} |C_j|^{\beta}\right] < 1$. In addition, suppose that

$$(Q, N, C_1, C_2, \dots) \stackrel{\mathcal{D}}{=} (-Q, N, C_1, C_2, \dots).$$
 (15)

Then,

$$P(|R_L| > t) \ge \frac{1}{2} P\left(\max_{\mathbf{i} \in \mathcal{T}} |\Pi_{\mathbf{i}} Q_{\mathbf{i}}| > t\right).$$

Remarks 4.2. (a) We call a weighted branching tree whose root vector satisfies (15) a symmetric tree. We will show how one can easily construct such trees in the proof of Corollary 4.4 below. (b) That the solution R_L for symmetric trees follows a power-law behavior with strictly positive constant of proportionality if and only if $Q \neq 0$ immediately follows from the preceding proposition and Theorem 3.4 applied to the weighted branching tree having root vector ($|Q|, N, |C_1|, |C_2|, ...$). (c) The proof of Proposition 4.1 follows the ideas of the Lévy-type maximal inequalities from [20] (see Theorem 1.1.1) adapted to weighted branching trees.

Proof of Proposition 4.1. We start by defining the process

$$W_0 = Q, \qquad W_k = \sum_{\mathbf{i} \in A_k} \prod_{\mathbf{i}} Q_{\mathbf{i}}, \quad k \in \mathbb{N},$$

and with some abuse of notation, the process

$$V_0 = |Q|, \qquad V_k = \bigvee_{\mathbf{i} \in A_k} |\Pi_{\mathbf{i}} Q_{\mathbf{i}}|, \quad k \in \mathbb{N}.$$

Next, consider the events

$$B_0 = \{V_0 > t\}, \qquad B_k = \left\{\max_{0 \le i \le k-1} V_i \le t, \ V_k > t\right\}, \quad k \in \mathbb{N},$$

and note that they are disjoint and satisfy

$$P\left(\max_{k\geq 0} V_k > t\right) = \sum_{k=0}^{\infty} P(B_k) = \sum_{k=0}^{\infty} \left(P\left(B_k, |R_L| > t\right) + P\left(B_k, |R_L| \le t\right)\right).$$
(16)

To analyze the second probability on the right hand side, for each $k = 1, 2, 3, \ldots$, let $m_k : \mathbb{N} \to \mathbb{N}^k$ be a bijective function, and use it to define the events $B_{k,1} = B_k \cap \{ |\Pi_{m_k(1)}Q_{m_k(1)}| > t \}$ and

$$B_{k,j} = B_k \cap \left\{ \max_{1 \le r \le j-1} |\Pi_{m_k(r)} Q_{m_k(r)}| \le t, \ |\Pi_{m_k(j)} Q_{m_k(j)}| > t \right\}, \quad j = 2, 3, 4, \dots,$$

where the convention is to set $\Pi_{m_k(r)} \equiv 0$ if $m_k(r) \notin A_k$. Note that the $\{B_{k,j}\}$ are disjoint and $P(B_k) = \sum_{j=1}^{\infty} P(B_{k,j})$. The key observation is that under the symmetry assumptions of the lemma we have that for any $k \geq 0$, and $r \in \mathbb{N}$ the sequences

$$\{(Q_{\mathbf{i}}, N_{\mathbf{i}}, C_{(\mathbf{i},1)}, C_{(\mathbf{i},2)}, \dots) : \mathbf{i} \in \mathcal{T}\}$$
and
$$\{(Q_{\mathbf{i}}, N_{\mathbf{i}}, C_{(\mathbf{i},1)}, C_{(\mathbf{i},2)}, \dots) : \mathbf{i} \in \mathcal{T}, \mathbf{i} \neq m_{k}(r)\}$$
$$\cup \{(-Q_{m_{k}(r)}, N_{m_{k}(r)}, C_{(m_{k}(r),1)}, C_{(m_{k}(r),2)}, \dots)\}$$

have the same distribution. It follows that for any $k \ge 0$ and $r \in \mathbb{N}$,

$$R_{L} = \sum_{j \neq k} W_{j} + \sum_{\mathbf{i} \in A_{k}, \mathbf{i} \neq m_{k}(r)} \Pi_{\mathbf{i}} Q_{\mathbf{i}} + \Pi_{m_{k}(r)} Q_{m_{k}(r)}$$
$$\stackrel{\mathcal{D}}{=} \sum_{j \neq k} W_{j} + \sum_{\mathbf{i} \in A_{k}, \mathbf{i} \neq m_{k}(r)} \Pi_{\mathbf{i}} Q_{\mathbf{i}} - \Pi_{m_{k}(r)} Q_{m_{k}(r)}$$
$$= R_{L} - 2\Pi_{m_{k}(r)} Q_{m_{k}(r)},$$

and since the events $\{B_{k,j}\}$ are insensitive to changes in the sign of the $\{Q_i\}$, we have that for any $k \ge 0$,

$$P(B_{k}, |R_{L}| \leq t) = \sum_{r=1}^{\infty} P(B_{k,r}, |R_{L}| \leq t)$$

$$= \sum_{r=1}^{\infty} P(B_{k,r}, |2\Pi_{m_{k}(r)}Q_{m_{k}(r)} - R_{L}| \leq t)$$

$$\leq \sum_{r=1}^{\infty} P(B_{k,r}, 2 |\Pi_{m_{k}(r)}Q_{m_{k}(r)}| - |R_{L}| \leq t)$$

$$= \sum_{r=1}^{\infty} P(B_{k,r}, |R_{L}| \geq 2 |\Pi_{m_{k}(r)}Q_{m_{k}(r)}| - t)$$

$$\leq \sum_{r=1}^{\infty} P(B_{k,r}, |R_{L}| > t) \quad (\text{since } |\Pi_{m_{k}(r)}Q_{m_{k}(r)}| > t \text{ on } B_{k,r})$$

$$= P(B_{k}, |R_{L}| > t).$$

This and equation (16) complete the proof.

We now proceed to the second step of our proof, the one that shows how to lower bound the endogenous solution R_L in the general case with the solution to the linear equation (13) on a closely related symmetric tree. The following technical lemma will be useful to explain our construction.

Lemma 4.3. Let X be a real-valued random variable and let $\mathbf{Y} = (Y_1, Y_2, ...) \in \mathbb{R}^{\infty}$ be a random vector on the same probability space. Define for $x \in \mathbb{R}$,

$$F_{\mathbf{Y}}(x) = E\left[1(X \le x) | \mathbf{Y}\right].$$

Then, $F_{\mathbf{Y}}(x)$ is nondecreasing in x a.s. and $F_{\mathbf{Y}}^{-1}(t) = \inf\{x \in \mathbb{R} : F_{\mathbf{Y}}(x) \ge t\}$ exists a.s. Moreover, if U_1, U_2 are two independent Uniform(0,1) random variables, independent of \mathbf{Y} , then

$$X_1 = F_{\mathbf{Y}}^{-1}(U_1)$$
 and $X_2 = F_{\mathbf{Y}}^{-1}(U_2)$

are identically distributed and conditionally independent given \mathbf{Y} , and

$$(X_i, Y_1, Y_2, \dots) \stackrel{\mathcal{D}}{=} (X, Y_1, Y_2, \dots), \qquad i = 1, 2.$$

Proof. That $F_{\mathbf{Y}}(x)$ is nondecreasing in x a.s. follows from the fact that $1(X \leq x)$ is nondecreasing. Moreover, the pseudo inverse $F_{\mathbf{Y}}^{-1}(t)$ is well defined for all $t \in \mathbb{R}$ and satisfies $F_{\mathbf{Y}}^{-1}(t) \leq x$ if and only if $t \leq F_{\mathbf{Y}}(x)$ for all t, x. Now consider the two random variables X_1 and X_2 from the statement of the lemma. Then, for any $x_1, x_2 \in \mathbb{R}$,

$$E [1(X_{1} \le x_{1}, X_{2} \le x_{2}) | \mathbf{Y}]$$

= $E [1(F_{\mathbf{Y}}^{-1}(U_{1}) \le x_{1})1(F_{\mathbf{Y}}^{-1}(U_{2}) \le x_{2}) | \mathbf{Y}]$
= $E [1(F_{\mathbf{Y}}^{-1}(U_{1}) \le x_{1}) | \mathbf{Y}] E [1(F_{\mathbf{Y}}^{-1}(U_{2}) \le x_{2}) | \mathbf{Y}]$
(since U_{1}, U_{2} and \mathbf{Y} are independent)
= $E [1(X_{1} \le x_{1}) | \mathbf{Y}] E [1(X_{2} \le x_{2}) | \mathbf{Y}],$

which shows the conditional independence given **Y**. Furthermore, for any $x \in \mathbb{R}$,

$$E[1(X_i \le x) | \mathbf{Y}] = E[1(F_{\mathbf{Y}}^{-1}(U_i) \le x) | \mathbf{Y}]$$

= $E[1(U_i \le F_{\mathbf{Y}}(x)) | \mathbf{Y}] = F_{\mathbf{Y}}(x) = E[1(X \le x) | \mathbf{Y}]$

for i = 1, 2. Hence, X_1 and X_2 have the same conditional distribution as $X|\mathbf{Y}$, from where it follows that for any measurable $A \subseteq \mathbb{R}^{\infty}$ and i = 1, 2,

$$P((X_i, Y_1, Y_2, \dots) \in A) = E[E[1((X_i, Y_1, Y_2, \dots) \in A) | \mathbf{Y}]]$$

= $E[E[1((X, Y_1, Y_2, \dots) \in A) | \mathbf{Y}]]$
= $P((X, Y_1, Y_2, \dots) \in A).$

We now derive as a corollary the strict positivity of the constant in Theorem 4.6 of [3] for the general case. Note that the result holds under no additional assumptions beyond those required in that theorem, in other words, we only require $E\left[\left(\sum_{i=1}^{N}|C_i|\right)^{\alpha}\right] < \infty$ when $\alpha > 1$ and $E\left[\left(\sum_{i=1}^{N}|C_i|^{\alpha/(1+\epsilon)}\right)^{1+\epsilon}\right] < \infty$ for some $\epsilon > 0$ (which does not imply $\rho_{\alpha+\epsilon} < \infty$) when $0 < \alpha \leq 1$; compare these to conditions (C) and (A) in [7], respectively. We also point out that Theorem 4.6 of [3] does not require the existence of the first root $0 < v < \alpha$ of the equation $E\left[\sum_{i=1}^{N}|C_i|^{v}\right] = 1$, but only the derivative condition $0 < E\left[\sum_{i=1}^{N}|C_i|^{\alpha}\log|C_i|\right] < \infty$. Finally, our main result on the asymptotic behavior of the minimal/endogenous solution R to the maximum equation (Theorem 3.4) gives that $H_L > 0$ provided Q is not a deterministic function of the weights $\{C_i\}$ and assuming all the other conditions in Theorem 4.6 of [3] are satisfied. We have included the entire statement of the theorem for completeness.

Corollary 4.4. Let $(Q, N, C_1, C_2, ...)$ be a random vector, with $N \in \mathbb{N} \cup \{\infty\}, \{C_i\}_{i \in \mathbb{N}}$ real-valued weights, Q a real-valued random variable with P(|Q| > 0) > 0 and R be the solution to (13) given by (14). Suppose that there exists $j \ge 1$ with $P(N \ge j, |C_j| > 0) > 0$ such that the measure $P(\log |C_j| \in du, |C_j| > 0, N \ge j)$ is non lattice, and that for some $\alpha > 0$, $E[|Q|^{\alpha}] < \infty$, $E\left[\sum_{i=1}^{N} |C_i|^{\alpha} \log |C_i|\right] > 0$ and $E\left[\sum_{i=1}^{N} |C_i|^{\alpha}\right] = 1$. In addition, assume

1.
$$E\left[\sum_{i=1}^{N} |C_i|\right] < 1$$
 and $E\left[\left(\sum_{i=1}^{N} |C_i|\right)^{\alpha}\right] < \infty$, if $\alpha > 1$; or,
2. $E\left[\left(\sum_{i=1}^{N} |C_i|^{\alpha/(1+\epsilon)}\right)^{1+\epsilon}\right] < \infty$ for some $0 < \epsilon < 1$, if $0 < \alpha \le 1$.

Then, provided Q is not a deterministic function of $(N, C_1, C_2, ...)$, we have

$$P(|R| > t) \sim H_L t^{-\alpha}, \qquad t \to \infty,$$

where $0 < H_L < \infty$ and

$$H_L = \frac{E\left[\left|\sum_{i=1}^N C_i R_i + Q\right|^{\alpha} - \sum_{i=1}^N |C_i R_i|^{\alpha}\right]}{\alpha E\left[\sum_{i=1}^N |C_i|^{\alpha} \log |C_i|\right]}.$$

Remark 4.5. Note that the same arguments used in the proof of Proposition 4.1 work, for any choice of $Q \neq 0$, if the weights $\{C_i\}$ are symmetric, in which case the strict positivity of H_L holds without any assumptions on Q.

Proof. Let $\{(Q_{\mathbf{i}}, N_{\mathbf{i}}, C_{(\mathbf{i},1)}, C_{(\mathbf{i},2)}, \dots)\}_{\mathbf{i} \in U}$ be a sequence of i.i.d. vectors and construct its corresponding random variable

$$R = \sum_{k=0}^{\infty} \sum_{\mathbf{i} \in A_k} \Pi_{\mathbf{i}} Q_{\mathbf{i}}.$$

Now use this sequence and Lemma 4.3 to construct a second i.i.d. sequence $\{(\hat{Q}_{\mathbf{i}}, N_{\mathbf{i}}, C_{(\mathbf{i},1)}, C_{(\mathbf{i},2)}, \dots)\}_{\mathbf{i} \in U}$ where

$$(\hat{Q}, N, C_1, C_2, \dots) \stackrel{\mathcal{D}}{=} (-Q, N, C_1, C_2, \dots)$$

and such that $\hat{Q}_{\mathbf{i}}$ and $Q_{\mathbf{i}}$ are conditionally independent given $(N_{\mathbf{i}}, C_{(\mathbf{i},1)}, C_{(\mathbf{i},2)}, \dots)$ for all $\mathbf{i} \in U$. Denote by \hat{R} the corresponding process

$$\hat{R} = \sum_{k=0}^{\infty} \sum_{\mathbf{i} \in A_k} \Pi_{\mathbf{i}} \hat{Q}_{\mathbf{i}},$$

~

and note that $|R| \stackrel{\mathcal{D}}{=} |\hat{R}|$. Next define

$$\overline{R} = \frac{R + \hat{R}}{2} = \sum_{k=0}^{\infty} \sum_{\mathbf{i} \in A_k} \prod_{\mathbf{i}} \left(\frac{Q_{\mathbf{i}} + \hat{Q}_{\mathbf{i}}}{2} \right) \triangleq \sum_{k=0}^{\infty} \sum_{\mathbf{i} \in A_k} \prod_{\mathbf{i}} \overline{Q}_{\mathbf{i}},$$

and observe that \overline{R} satisfies the conditions of Proposition 4.1, and therefore

$$P(|\overline{R}| > t) \ge \frac{1}{2} P\left(\max_{\mathbf{i} \in \mathcal{T}} |\Pi_{\mathbf{i}} \overline{Q}_{\mathbf{i}}| > t\right).$$

Note that the assumption that Q is not a deterministic function of $(N, C_1, C_2, ...)$ implies that $\overline{Q} \neq 0$. Moreover, by Theorem 3.4,

$$P\left(\max_{\mathbf{i}\in\mathcal{T}}|\Pi_{\mathbf{i}}\overline{Q}_{\mathbf{i}}|>t\right)\sim\overline{H}t^{-\alpha}$$

as $t \to \infty$ for some constant $0 < \overline{H} < \infty$. The last step is to note that

$$P(|\overline{R}| > t) \le P\left(|R| + |\hat{R}| > 2t\right) \le P(|R| > t) + P(|\hat{R}| > t) = 2P(|R| > t).$$

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- P.R. Jelenković, M. Olvera-Cravioto, Information ranking and power laws on trees, Adv. Appl. Prob. 42 (4) (2010) 1057–1093.
- [2] P.R. Jelenković, M. Olvera-Cravioto, Implicit renewal theory and power tails on trees, Adv. Appl. Prob. 44 (2) (2012) 528–561.
- [3] P.R. Jelenković, M. Olvera-Cravioto, Implicit renewal theory for trees with general weights, Stochastic Process. Appl. 122 (9) (2012) 3209– 3238.
- [4] G. Alsmeyer, M. Meiners, Fixed points of inhomogeneous smoothing transforms, J. Diff. Equations Appl. 18 (8) (2012) 1287–1304.
- G. Alsmeyer, M. Meiners, Fixed points of the smoothing transform: Two-sided solutions, Probab. Theory Relat. Fields 155 (1-2) (2013) 165– 199.
- [6] M. Mirek, On fixed points of a generalized multidimensional affine recursion, Probab. Theory Relat. Fields 156 (2013) 665–705.
- [7] G. Alsmeyer, E. Damek, S. Mentemeier, Precise tail index of fixed points of the two-sided smoothing transform, in: G. A. M. Löwe (Ed.), Random Matrices and Iterated Random Functions, Springer Proceedings in Mathematics & Statistics, Münster, Germany, 2011.
- [8] D. Buraczewski, E. Damek, S. Mentemeier, M. Mirek, Heavy tailed solutions of multivariate smoothing transforms, arXiv:1206.1709 (2013) 1–35.
- [9] U. Rösler, L. Rüschendorf, The contraction method for recursive algorithms, Algorithmica 29 (1-2) (2001) 3–33.
- [10] R. Neininger, L. Rüschendorf, A general limit theorem for recursive algorithms and combinatorial structures, Ann. Appl. Probab. 14 (1) (2004) 378–418.
- [11] J.A. Fill, S. Janson, Approximating the limiting Quicksort distribution, Random Structures Algorithms 19 (3-4) (2001) 376–406.

- [12] Y. Volkovich, N. Litvak, Asymptotic analysis for personalized web search, Adv. Appl. Prob. 42 (2) (2010) 577–604.
- [13] D.J. Aldous, A. Bandyopadhyay, A survey of max-type recursive distributional equation, Annals of Applied Probability 15 (2) (2005) 1047– 1110.
- [14] C.M. Goldie, Implicit renewal theory and tails of solutions of random equations, Ann. Appl. Probab. 1 (1) (1991) 126–166.
- [15] F.I. Karpelevich, M.Ya. Kelbert, Yu.M. Suhov, Higher-order Lindley equations, Stochastic Process. Appl. 53 (1994) 65–96.
- [16] J.D. Biggins, Lindley-type equations in the branching random walk, Stochastic Process. Appl. 75 (1998) 105–133.
- [17] G. Alsmeyer, U. Rösler, A stochastic fixed point equation for weighted minima and maxima, Annales de l'Institut Henri Poincaré 44 (1) (2008) 89–103.
- [18] P. Jagers, U. Rösler, Mathematics and Computer Science III, Birkhäuser, Basel, 2004, Ch. Stochastic fixed points involving the maximum, pp. 325–338.
- [19] P. Billingsley, Probability and Measure, 3rd Edition, Wiley-Interscience, New York, 1995.
- [20] V.H. de la Peña, E. Giné, Decoupling: From Dependence to Independence, Probability and its Applications, Springer, New York, 1999.